

this continuous rvs;  
 time: joint, marginal &  
 conditional distributions

Read: Dsch.  
 3,4

AMS 131  
 9 Aug 17  
 (doc. ①  
 com.  
 notes)

Case  
 Study

next functions of rvs;  
 time: expected value

Re  
 Poisson process (D.S.  
 p. 294)

You work for a bank that has lately been  
 receiving complaints about long waiting times  
 in ~~the~~ bank teller lines. To quantify the  
 extent of the problem you gather a dataset

time	event
9 AM	bank opens
9:01 am	arrival
9:03 am	arrival
9:04 am	arrival
9:05 am	departure
⋮	⋮

like the one at left, on  
 Mon, Tue, ..., Sat in a  
 randomly chosen week. This  
 is a new type of dataset  
 for us: it unfolds in time.  
 There are two equivalent ways

to keep track of the arrivals: you could  
 let  $N(t) = \# \text{arrivals in } [0, t]$  or you

could keep track of the inter-arrival times (times between arrivals)  $T_1, T_2, \dots$

Here we'll look at  $N(t)$ . Definition:

A stochastic process is a collection of random variables indexed by elements of an indexing set, usually  $t \in T$ , where  $t$  represents time.

$N(t), t \in [0^{\text{AM}}, t_{\text{max}}]$  is an example of (eg.  $\uparrow$   $\text{Spn}$ )

a stochastic process, in which the time index is continuous;  $\{T_1, T_2, \dots\}$  (interarrival times)

is an example of a stochastic process with discrete time index.  $\uparrow$   
 $(T_t, t=1, 2, \dots)$

What may reasonably be assumed about the random behavior of  $N(t)$ ?

Assumption 1 | The numbers of arrivals (nonoverlapping) <sup>(3)</sup>  
in any collection of disjoint time intervals  
are (mutually) independent (this is reasonable  
if unrelated customers arrive at the  
bank haphazardly in time).

Assumption 2

To more short  
 $P$ (arrivals in time interval  $[s, s+t]$ ,  $t$  small)  
is proportional to  $t$ , for example  $\lambda \cdot t$  for  
a rate parameter  $\lambda > 0$  (this is reasonable  
if the arrivals process is smooth rather  
than lumpy).

Definition | To say

that a function  $f(t)$  of  $t$  (small)  
is  $o(t)$   
(read little-oh of  $t$ ) is to say that

$$\lim_{t \rightarrow 0} \frac{f(t)}{t} = 0$$

In other words,  $f(t)$  approaches  
0 as  $t \rightarrow 0$  at a rate faster  
than  $t$  itself.

The mathematically formal way to state ④

Assumption 2 is then

$$P(\text{at least one arrival in } [s, s+t]) = \lambda t + o(t) \quad (t \text{ small})$$

independent of s

Assumption 3 } <sup>nearly</sup> (simultaneous arrivals are rare)

$$P(\text{2 or more arrivals in } [s, s+t]) = o(t) \quad (t \text{ small})$$

Remarkably, these 3 simple & often plausible assumptions specify the probability behavior of  $N(t)$  uniquely (see DS exercise 16, p. 296) for a proof of the following result).

$\lambda$  constant  $\leftrightarrow N(t)$  is a stationary stochastic process

of a day (eg. 10am - 11.30am)

Note that Assumption 2 implies that the rate parameter  $\lambda$  is constant in time; this would be unrealistic in the bank problem over an entire day but would be reasonable during stable subsets

Definition For any  $\lambda > 0$ , a random variable  $I$  has the Poisson distribution

with parameter  $\lambda$ , if its pf is  $(\text{Poisson}(\lambda))$  S. Poisson (1805-1850)

$$f(y | \lambda) = P(I = y | \lambda) = \begin{cases} \frac{\lambda^y e^{-\lambda}}{y!} & \text{for } y = 0, 1, \dots \\ 0 & \text{else} \end{cases}$$

(first person to study this was Abraham de Moivre (1711))

Result Under Assumptions

1-3 above, if  $I = (\# \text{ of arrivals in any time interval of length } t)$

then  $I \sim \text{Poisson}(\lambda t)$ .

Definition

A Poisson process with rate  $\lambda$  per unit

time is a stochastic process satisfying for any  $s > 0$

- (1) # arrivals in  $[s, s+t) \sim \text{Poisson}(\lambda t)$
- (2) # of arrivals in disjoint time intervals are independent

Restatement  
of Result

Under Assumptions 1-3 (6)

above,  $N(t)$  is a Poisson  
process with rate parameter  $\lambda$ .

Exploring the Poisson distribution

$$f(y|\lambda) = P(\bar{I} = y | \lambda) = \begin{cases} \frac{\lambda^y e^{-\lambda}}{y!} & y = 0, 1, \dots \\ 0 & \text{else} \end{cases}$$

(for  $\lambda > 0$ ).

This is our

first example of a discrete rv that takes  
on a countably infinite number of  
possible values.

Comment:

In reality  
infinite

we don't expect a Poisson rv to be <sup>infinite</sup> or  
"nearly infinite"; all the statement

( $y = 0, 1, \dots$ ) means is that ahead of time  
we can't place a <sup>fixed</sup> upper bound on  $\bar{I}$ .